Modeling Group fMRI Data

NITP 2011

Overview

- What is a mixed effects model
  - Fixed effects
  - Random effects
- 2-stage summary statistics approach
- How do different software packages work?
- Overview FSL modeling options

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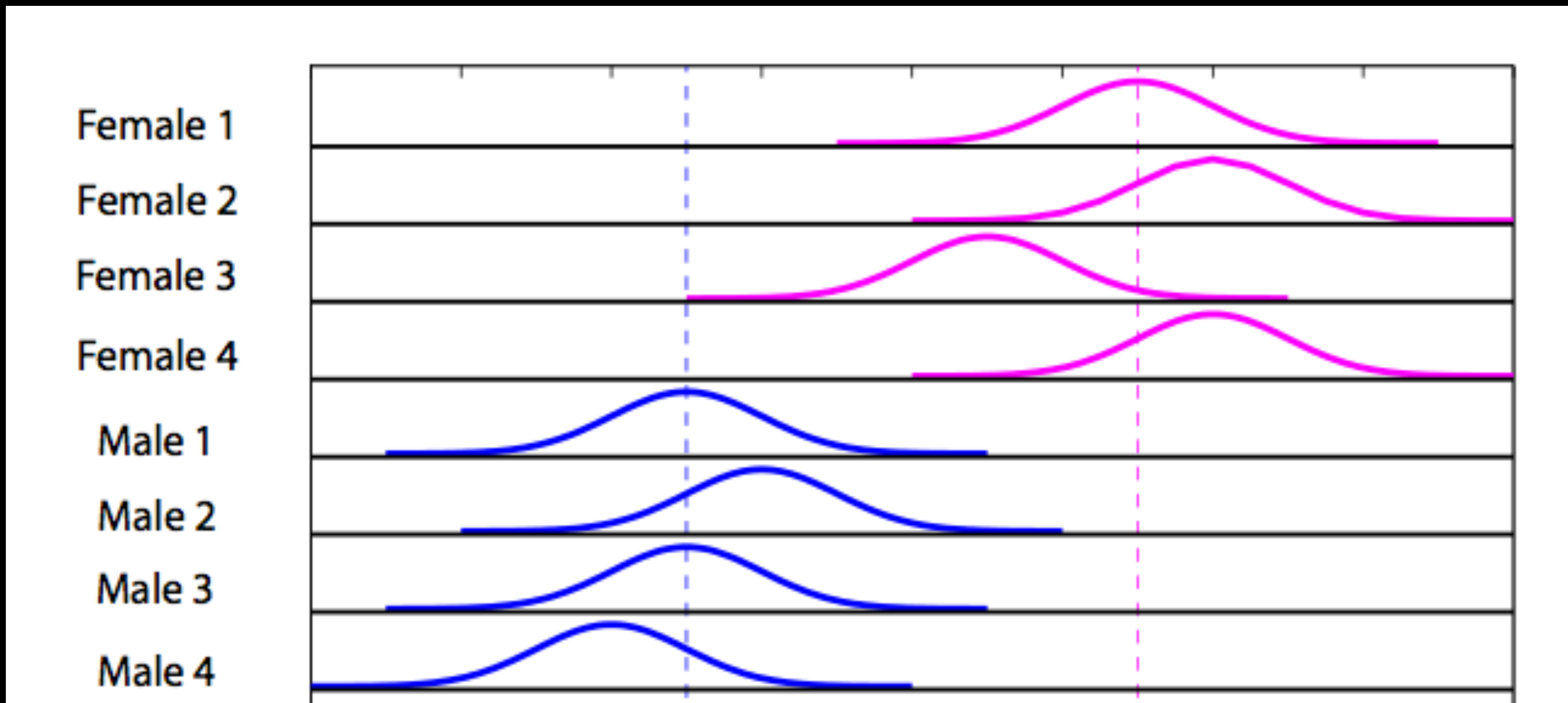
Mixed Model Motivation

- Start with a simple ANOVA example
- Study: Is hair length different between males and females?

Start: 1 hair per person

- Two sources of variability
  - Variance of hair length within person
  - Variance of hair length between people

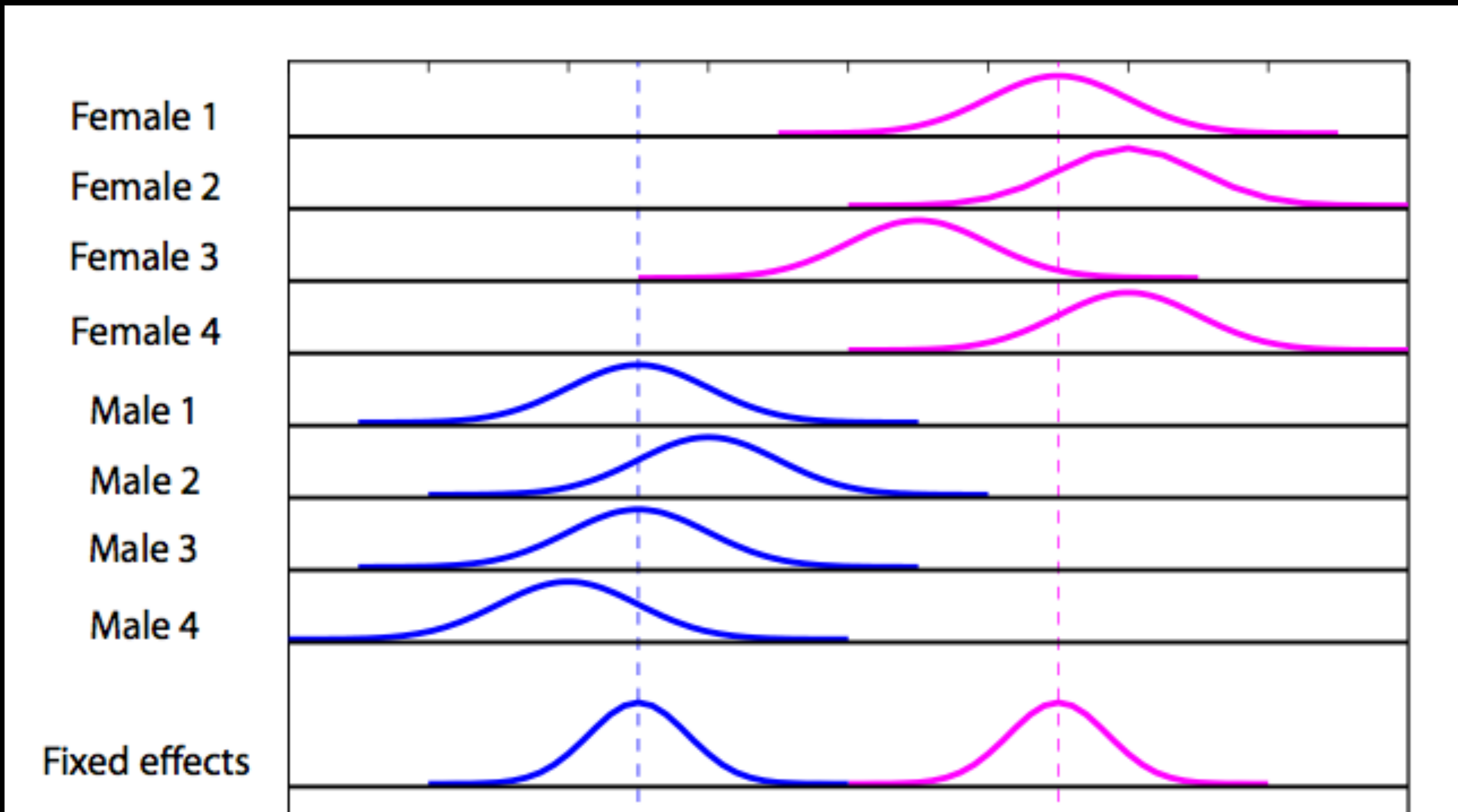
- Assume within-subject variance is 1



Each distribution has a variance of 1

Fixed effects analysis

- We're only interested in these exact 4 men and 4 women
  - $\sigma^2_{\mathrm{FFX}} = \frac{1}{4}\sigma^2_{\mathrm{W}} = 0.25$

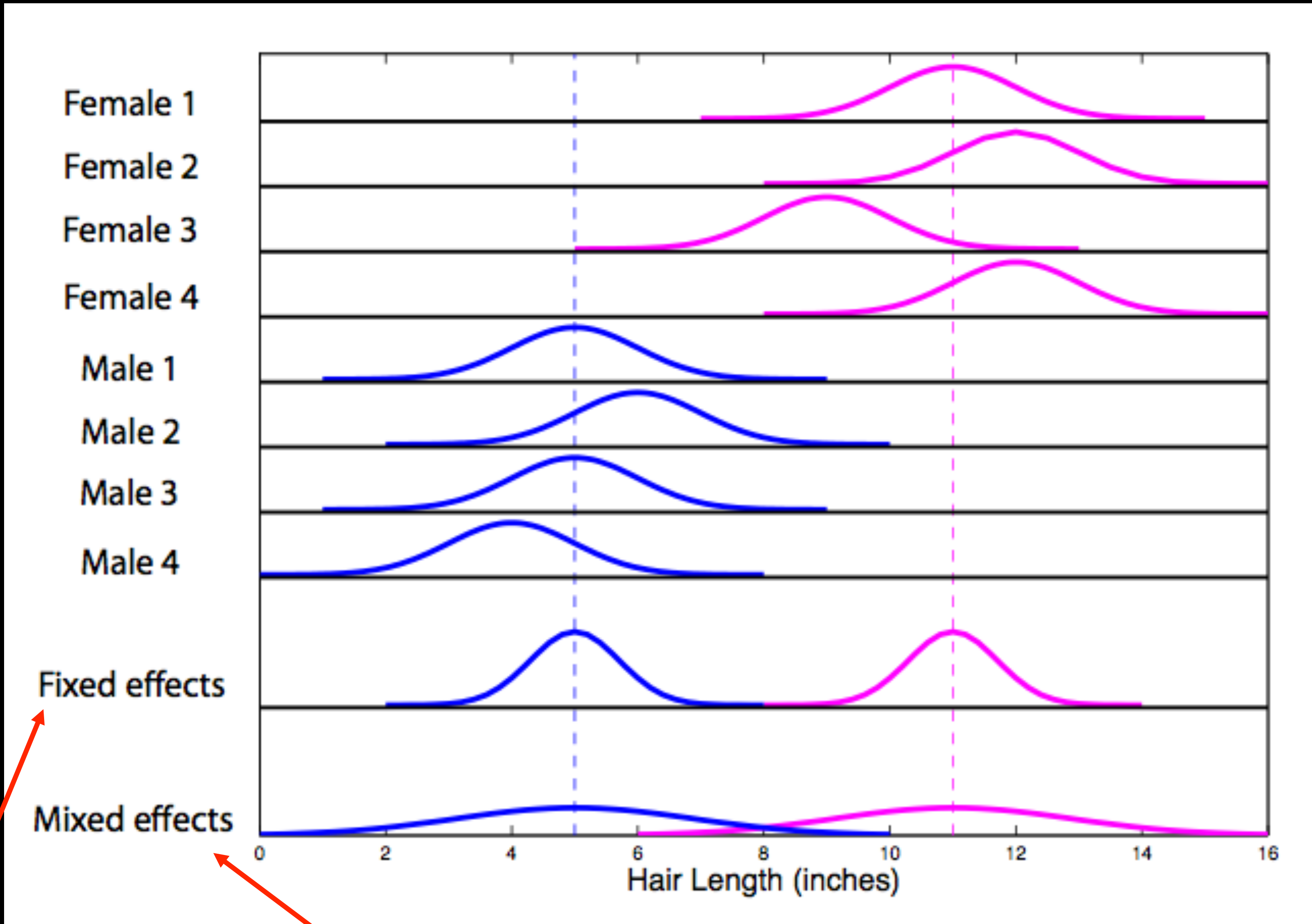


$$\sigma^2_{\mathrm{FFX}} = \frac{1}{4}\sigma^2_{\mathrm{W}} = 0.25$$

Mixed effects

- Include both within *and* between subject variances
- Adding a between subject means subject is random
  - Anything with a variance is random!

$$\sigma^2_{\mathrm{MFX}} = \sigma^2_{\mathrm{W}}/4 + \sigma^2_{\mathrm{B}}/4 = 1/4 + 49/4 = 12.5$$



$$\sigma^2_{\mathrm{FFX}} = \frac{1}{4}\sigma^2_{\mathrm{W}} = 0.25$$

$$\sigma^2_{\mathrm{MFX}} = \sigma^2_{\mathrm{W}}/4 + \sigma^2_{\mathrm{B}}/4 = 1/4 + 49/4 = 12.5$$

Multiple hairs per subject

- Fixed effects variance
  - $\sigma^2_{\text{FFX}} = \frac{1}{4}\sigma^2_{\text{W}}/25 = 0.01$
- Mixed effects variance
  - $\sigma^2_{\text{MFX}} = \frac{1}{4}\sigma^2_{\text{W}}/25 + \frac{1}{4}\sigma^2_{\text{B}} = 12.26$

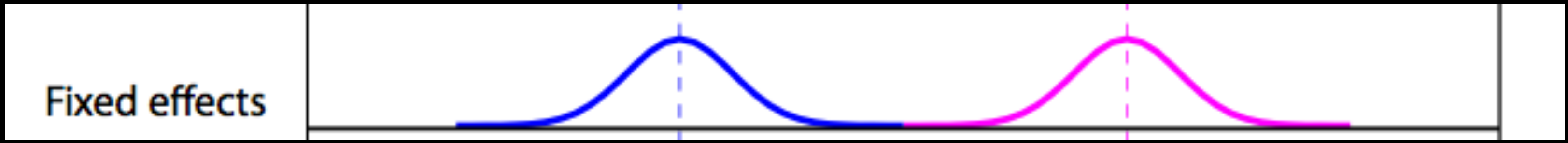
Multiple hairs per subject

- Fixed effects variance
  - $\sigma^2_{\text{FFX}} = \frac{1}{4}\sigma^2_{\text{W}}/25 = 0.01$
- Mixed effects variance
  - $\sigma^2_{\text{MFX}} = \frac{1}{4}\sigma^2_{\text{W}}/25 + \frac{1}{4}\sigma^2_{\text{B}} = 12.26$

Between subject variance typically dominates

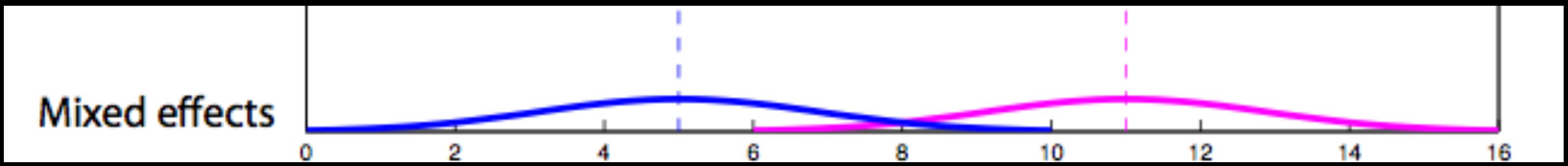
Wrong model leads to wrong conclusion

- Scenario 1: Fixed effects model
  - Significant difference in hair length
  - Result only applies to these 8 subjects



Wrong model leads to wrong conclusion

- Scenario 2: Mixed effects model
  - Cannot conclude there is a difference in hair length

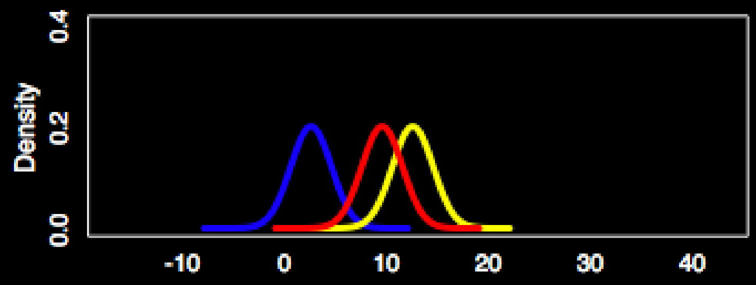


Mixed Model Comments

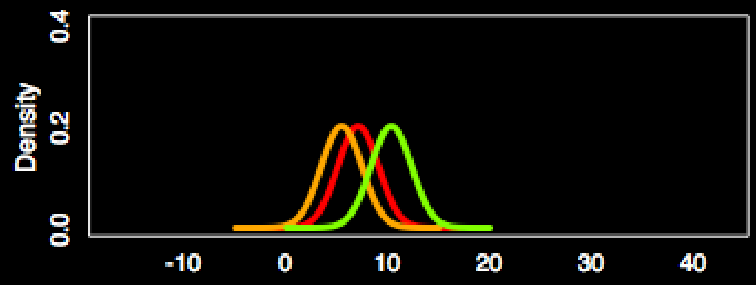
- If you fail to include a random effect when there is one
  - Results only apply to that data sample
  - P-values are smaller than mixed model p-values

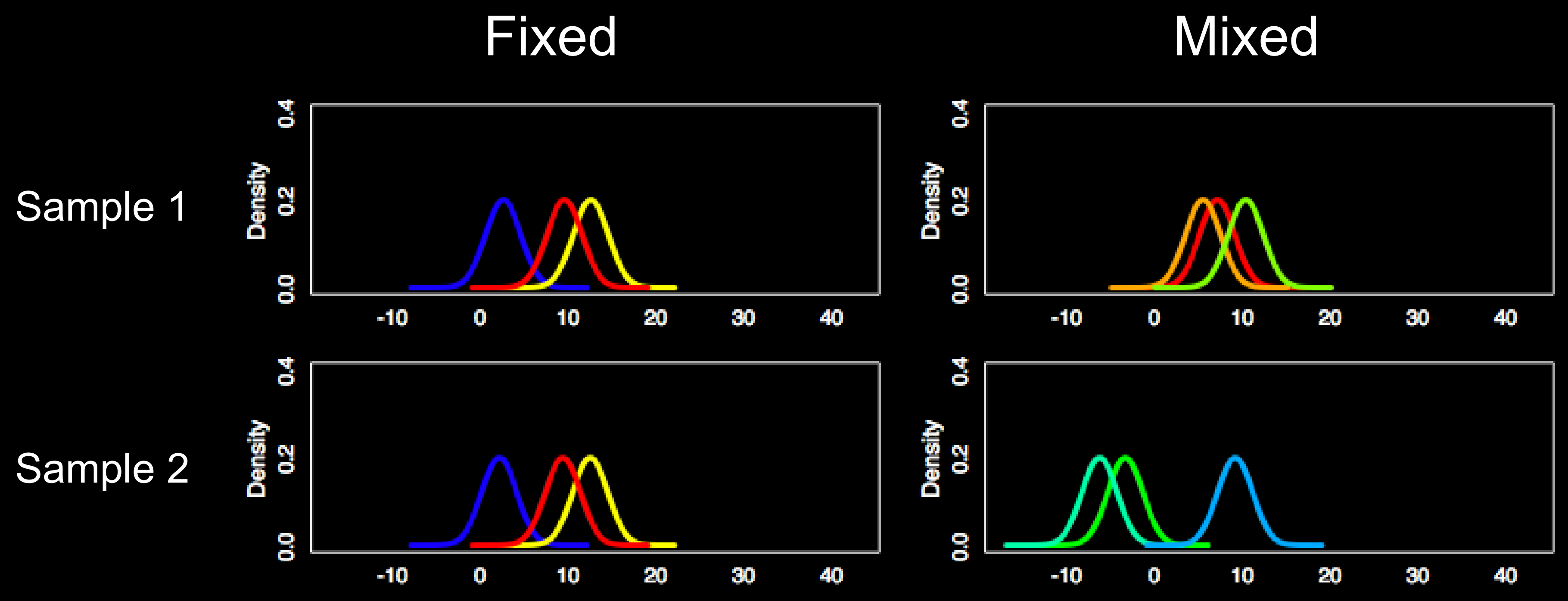
Sample 1

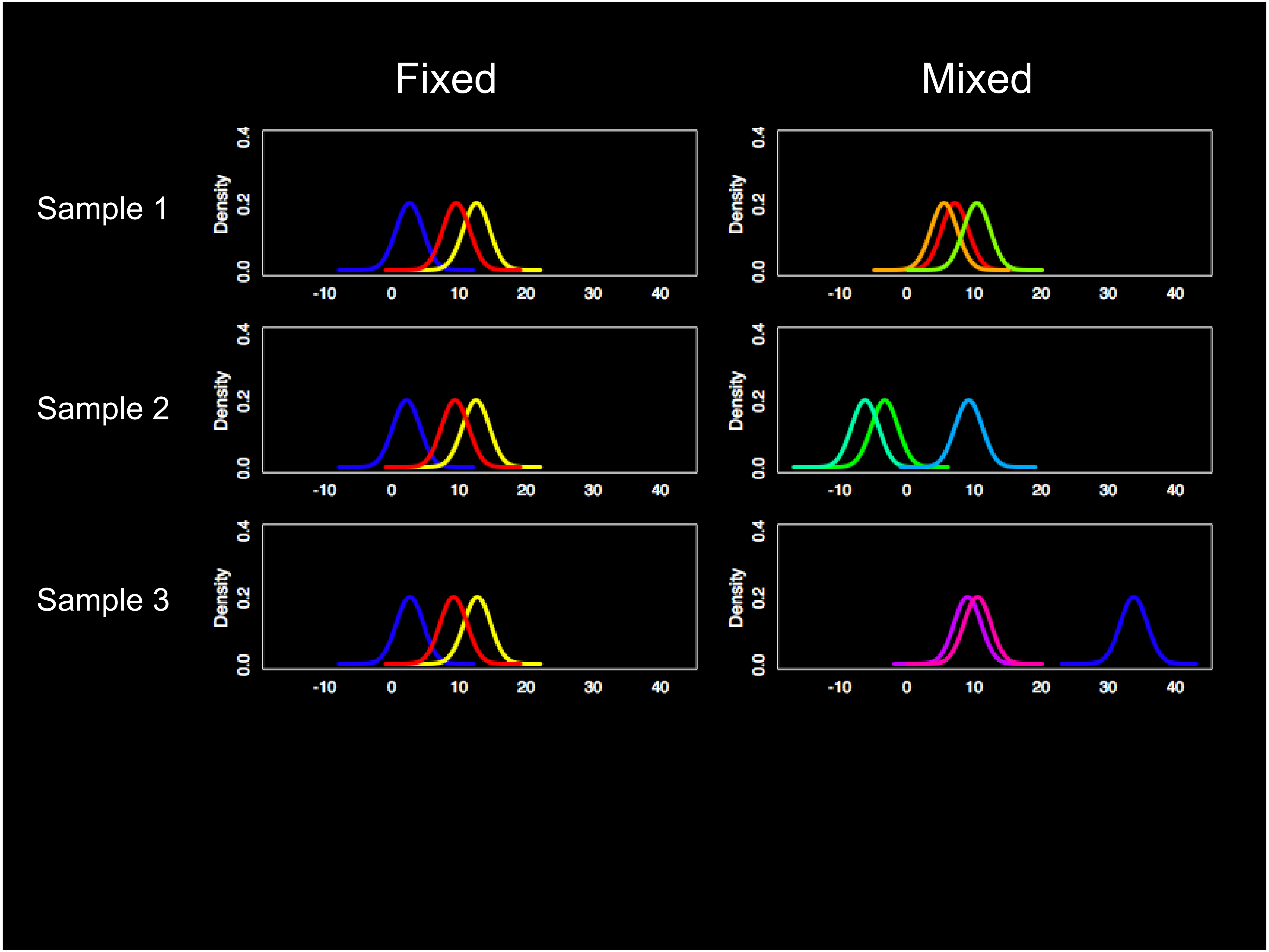
Fixed



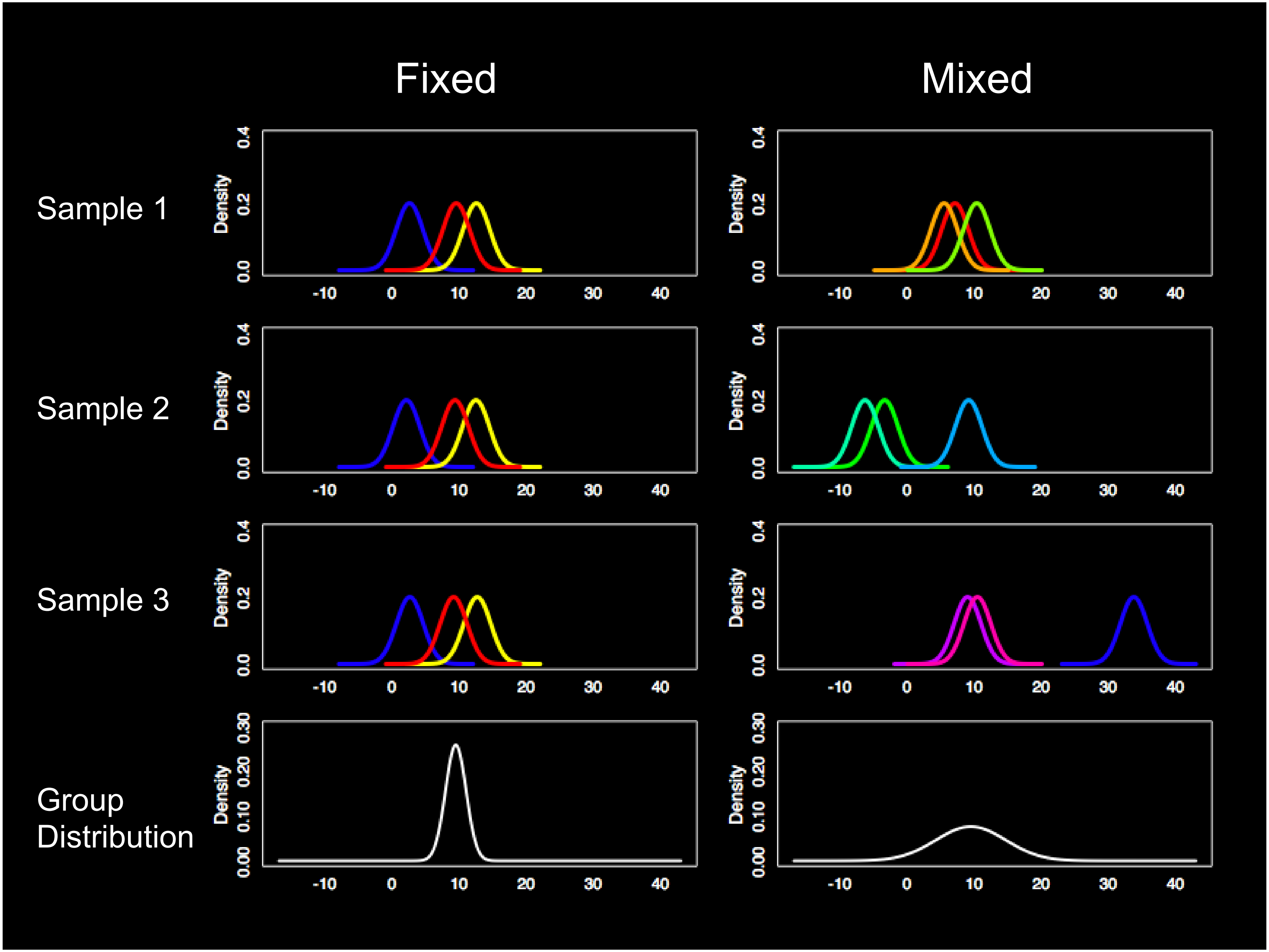
Mixed




Fixed
Mixed
Sample 1
Sample 2
Density
0.0
0.2
0.4
-10
0
10
20
30
40



Fixed
Mixed
Sample 1
Sample 2
Sample 3
Density
0.0
0.2
0.4
-10
0
10
20
30
40


Fixed
Mixed
Sample 1
Sample 2
Sample 3
Group
Distribution
Density
0.0
0.2
0.4
0.00
0.10
0.20
0.30
-10
0
10
20
30
40

Important points so far…

- Ignoring random subject effect means that you' re ignoring the fact that these subjects were randomly sampled
  – Inference only applies to sample you collected
- Including random subject effect *always* increases your variance

Important points so far…

- What has a bigger impact in reducing variance?
  - Adding more hairs per subject?
  - Adding more subjects?

Important points so far…

- What has a bigger impact in reducing variance?
  - Adding more hairs per subject?
  - Adding more subjects?

1 hair per subject

$$\sigma^2_{\text{MFX}} = \sigma^2_{\text{W}}/4 + \sigma^2_{\text{B}}/4 = 12.5$$

25 hairs per subject

$$\sigma^2_{\text{MFX}} = \frac{1}{4}\sigma^2_{\text{W}}/25 + \frac{1}{4}\sigma^2_{\text{B}} = 12.26$$

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Fixed effects model:

modeling the mean of 3 females, 20 hairs

Fixed effect

Residual error

$$\begin{pmatrix} Y_{1,1} \\ \vdots \\ Y_{1,20} \\ Y_{2,1} \\ \vdots \\ Y_{3,20} \end{pmatrix} = \begin{pmatrix} 1 \\ 1 \\ 1 \\ 1 \\ \vdots \\ 1 \end{pmatrix} \beta_g + \begin{pmatrix} \epsilon_{1,1} \\ \epsilon_{1,2} \\ \epsilon_{1,3} \\ \epsilon_{2,1} \\ \vdots \\ \epsilon_{N,3} \end{pmatrix}, \quad \epsilon_{i,j} \sim N(0, \sigma^2_{win})$$

Mixed Effects Model

Stage 1
$$Y = X\beta + \epsilon$$

$$\begin{pmatrix} Y_{1,1} \\ \vdots \\ Y_{1,20} \\ Y_{2,1} \\ \vdots \\ Y_{3,20} \end{pmatrix} = \begin{pmatrix} 1 & 0 & 0 \\ \vdots & \vdots & \vdots \\ 1 & 0 & 0 \\ 0 & 1 & 0 \\ \vdots & \vdots & \vdots \\ 0 & 0 & 1 \end{pmatrix} \begin{pmatrix} \beta_1 \\ \beta_2 \\ \beta_3 \end{pmatrix} + \begin{pmatrix} \epsilon_{1,1} \\ \epsilon_{1,2} \\ \epsilon_{1,3} \\ \epsilon_{2,1} \\ \vdots \\ \epsilon_{N,3} \end{pmatrix}, \epsilon_{i,j} \sim N(0, \sigma^2_{win})$$

Stage 2
$$\beta = X_g\beta_g + \eta$$

Random effect

$$\begin{pmatrix} \beta_1 \\ \beta_2 \\ \beta_3 \end{pmatrix} = \begin{pmatrix} 1 \\ 1 \\ 1 \end{pmatrix} \beta_g + \begin{pmatrix} \eta_1 \\ \eta_2 \\ \eta_3 \end{pmatrix}, \quad \eta_i \sim N(0, \sigma^2_{btwn})$$

Mixed Effects Model

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Stage 2 $$\beta = X_g\beta_g + \eta$$

Random effect

$$\begin{pmatrix} \beta_1 \\ \beta_2 \\ \beta_3 \end{pmatrix} = \begin{pmatrix} 1 \\ 1 \\ 1 \end{pmatrix} \beta_g + \begin{pmatrix} \eta_1 \\ \eta_2 \\ \eta_3 \end{pmatrix}, \quad \eta_i \sim N(0, \sigma^2_{btwn})$$

Mixed Effects Model: All-In-One

$$Y = XX_g\beta_g + \quad X\eta \quad + \quad \epsilon$$

$$\begin{pmatrix} Y_{1,1} \\ \vdots \\ Y_{1,20} \\ Y_{2,1} \\ \vdots \\ Y_{3,20} \end{pmatrix} = \begin{pmatrix} 1 \\ 1 \\ 1 \\ 1 \\ \vdots \\ 1 \end{pmatrix} \beta_g + \underbrace{\begin{pmatrix} 1 & 0 & 0 \\ \vdots & \vdots & \vdots \\ 1 & 0 & 0 \\ 0 & 1 & 0 \\ \vdots & \vdots & \vdots \\ 0 & 0 & 1 \end{pmatrix} \begin{pmatrix} \eta_1 \\ \eta_2 \\ \eta_3 \end{pmatrix} + \begin{pmatrix} \epsilon_{1,1} \\ \epsilon_{1,2} \\ \epsilon_{1,3} \\ \epsilon_{2,1} \\ \vdots \\ \epsilon_{N,3} \end{pmatrix}}_{\text{Variance Terms}}$$

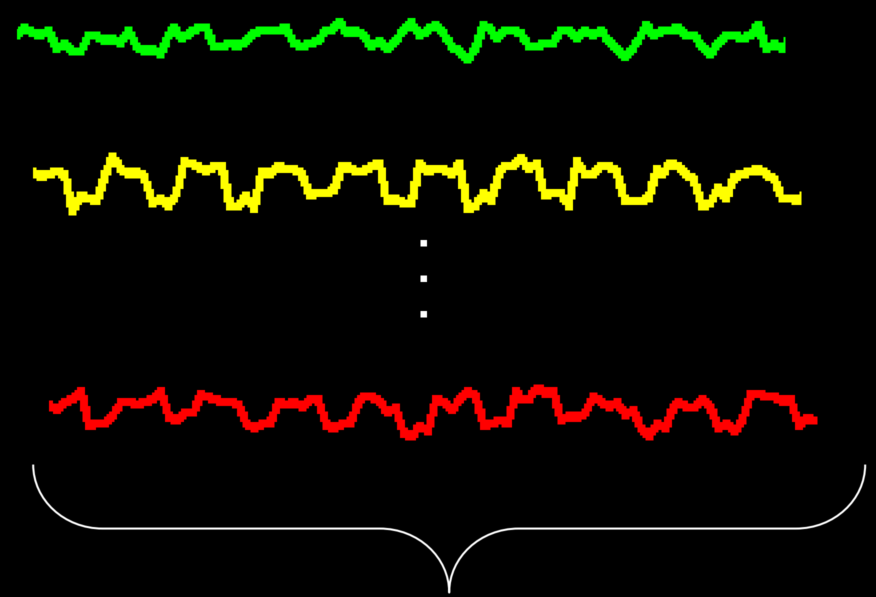
How does this relate to fMRI?

Subject 1

Subject 2

⋮

Subject N

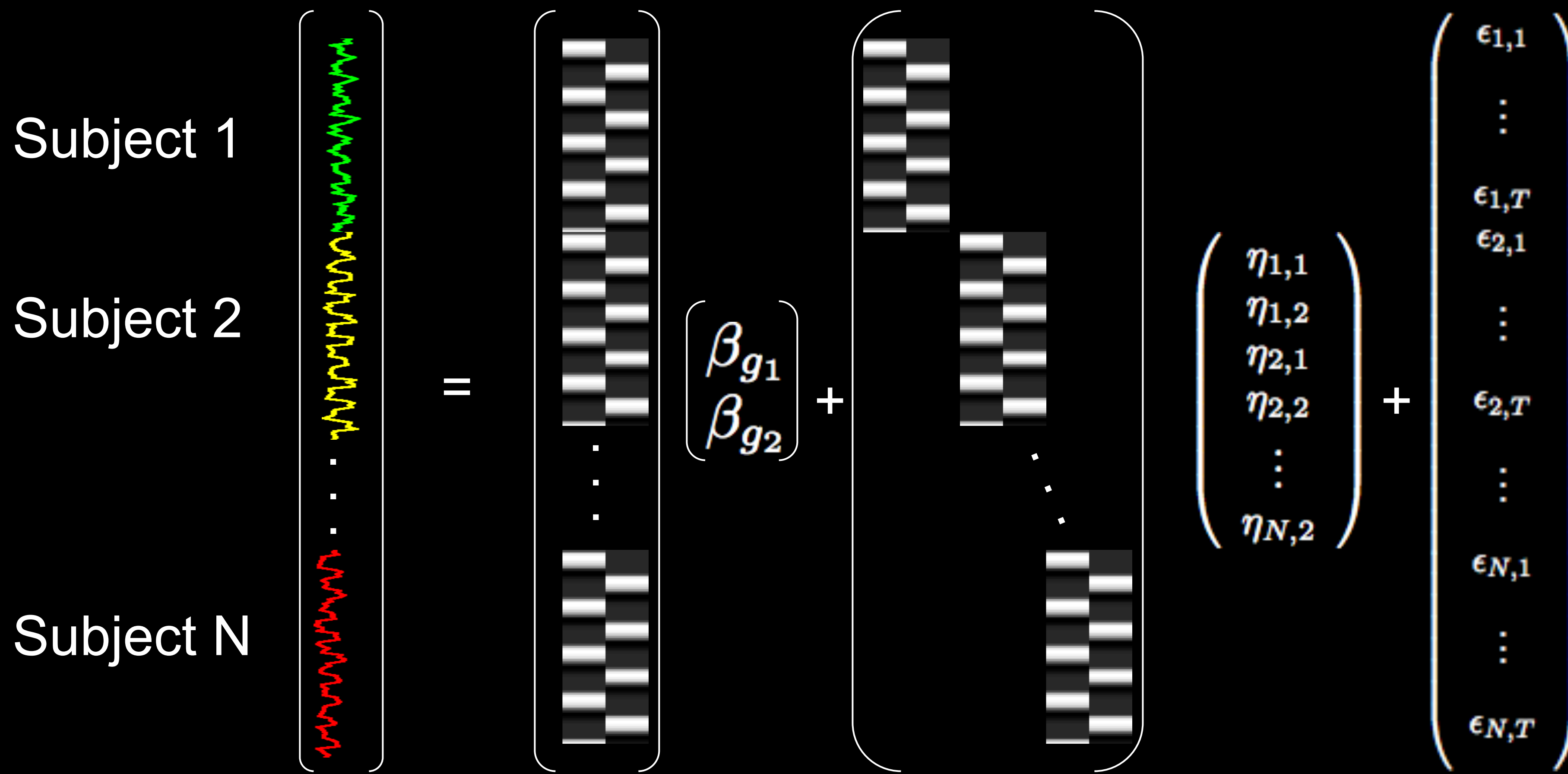


Each time series is a collection of data grouped by subject

A random subject effect is necessary to apply inference to total population

Mixed Model for fMRI Data

- fMRI data are more complicated than the hair length example
  - Not typically estimating an intercept
  - Time series are temporally autocorrelated
  - Time series can be quite long
- Let's take a look at the model!
  - A study with 2 stimuli of interest



$$\mathrm{Var}(\eta_{i,1}) = \sigma^2_{btwn_1}$$
$$\mathrm{Var}(\eta_{i,2}) = \sigma^2_{btwn_2}$$

$$\mathrm{Cov}\begin{pmatrix} \epsilon_{i,1} \\ \epsilon_{i,2} \\ \vdots \\ \epsilon_{i,T} \end{pmatrix} = \sigma^2_{win_i} V_i$$

Yuck!

- Computationally intensive
  - Large matrices that need to be inverted
- What if we add another subject?
  - Must estimate *whole* model for all subjects

Recall the two stages

Stage 1

$$Y = X\beta + \epsilon$$

$$\begin{pmatrix} Y_{1,1} \\ \vdots \\ Y_{1,20} \\ Y_{2,1} \\ \vdots \\ Y_{3,20} \end{pmatrix} = \begin{pmatrix} 1 & 0 & 0 \\ \vdots & \vdots & \vdots \\ 1 & 0 & 0 \\ 0 & 1 & 0 \\ \vdots & \vdots & \vdots \\ 0 & 0 & 1 \end{pmatrix} \begin{pmatrix} \beta_1 \\ \beta_2 \\ \beta_3 \end{pmatrix} + \begin{pmatrix} \epsilon_{1,1} \\ \epsilon_{1,2} \\ \epsilon_{1,3} \\ \epsilon_{2,1} \\ \vdots \\ \epsilon_{N,3} \end{pmatrix}, \epsilon_{i,j} \sim N(0, \sigma^2_{win})$$

Stage 2

$$\beta = X_g\beta_g + \eta$$

$$\begin{pmatrix} \beta_1 \\ \beta_2 \\ \beta_3 \end{pmatrix} = \begin{pmatrix} 1 \\ 1 \\ 1 \end{pmatrix} \beta_g + \begin{pmatrix} \eta_1 \\ \eta_2 \\ \eta_3 \end{pmatrix}, \quad \eta_i \sim N(0, \sigma^2_{btwn})$$

Two-Stage Summary Statistics

Stage 1

$$\begin{pmatrix} Y_{1,1} \\ \vdots \\ Y_{1,20} \end{pmatrix} = \begin{pmatrix} 1 \\ \vdots \\ 1 \end{pmatrix} \beta_1 + \begin{pmatrix} \epsilon_{1,1} \\ \vdots \\ \epsilon_{1,20} \end{pmatrix}$$

$$\begin{pmatrix} Y_{2,1} \\ \vdots \\ Y_{2,20} \end{pmatrix} = \begin{pmatrix} 1 \\ \vdots \\ 1 \end{pmatrix} \beta_2 + \begin{pmatrix} \epsilon_{2,1} \\ \vdots \\ \epsilon_{2,20} \end{pmatrix} \qquad \epsilon_{i,j} \sim N(0, \sigma^2_{win})$$

$$\begin{pmatrix} Y_{3,1} \\ \vdots \\ Y_{3,20} \end{pmatrix} = \begin{pmatrix} 1 \\ \vdots \\ 1 \end{pmatrix} \beta_3 + \begin{pmatrix} \epsilon_{3,1} \\ \vdots \\ \epsilon_{3,20} \end{pmatrix}$$

Two-Stage Summary Statistics

Stage 1

$$\begin{pmatrix} Y_{1,1} \\ \vdots \\ Y_{1,20} \end{pmatrix} = \begin{pmatrix} 1 \\ \vdots \\ 1 \end{pmatrix} \beta_1 + \begin{pmatrix} \epsilon_{1,1} \\ \vdots \\ \epsilon_{1,20} \end{pmatrix}$$

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Stage 2

Use first stage estimates

$$\begin{pmatrix} \hat{\beta}_1 \\ \hat{\beta}_2 \\ \hat{\beta}_3 \end{pmatrix} = \begin{pmatrix} 1 \\ 1 \\ 1 \end{pmatrix} \beta_g + \begin{pmatrix} \eta^*_1 \\ \eta^*_2 \\ \eta^*_3 \end{pmatrix}, \quad \mathrm{Var}(\eta^*_i) = \frac{\sigma^2_{win}}{W} + \sigma^2_{btwn}$$

Two-Stage Summary Statistics

Stage 1

$$\begin{pmatrix} Y_{1,1} \\ \vdots \\ Y_{1,20} \end{pmatrix} = \begin{pmatrix} 1 \\ \vdots \\ 1 \end{pmatrix} \beta_1 + \begin{pmatrix} \epsilon_{1,1} \\ \vdots \\ \epsilon_{1,20} \end{pmatrix}$$

$$\begin{pmatrix} Y_{2,1} \\ \vdots \\ Y_{2,20} \end{pmatrix} = \begin{pmatrix} 1 \\ \vdots \\ 1 \end{pmatrix} \beta_2 + \begin{pmatrix} \epsilon_{2,1} \\ \vdots \\ \epsilon_{2,20} \end{pmatrix} \qquad \epsilon_{i,j} \sim N(0, \sigma^2_{win})$$

$$\begin{pmatrix} Y_{3,1} \\ \vdots \\ Y_{3,20} \end{pmatrix} = \begin{pmatrix} 1 \\ \vdots \\ 1 \end{pmatrix} \beta_3 + \begin{pmatrix} \epsilon_{3,1} \\ \vdots \\ \epsilon_{3,20} \end{pmatrix}$$

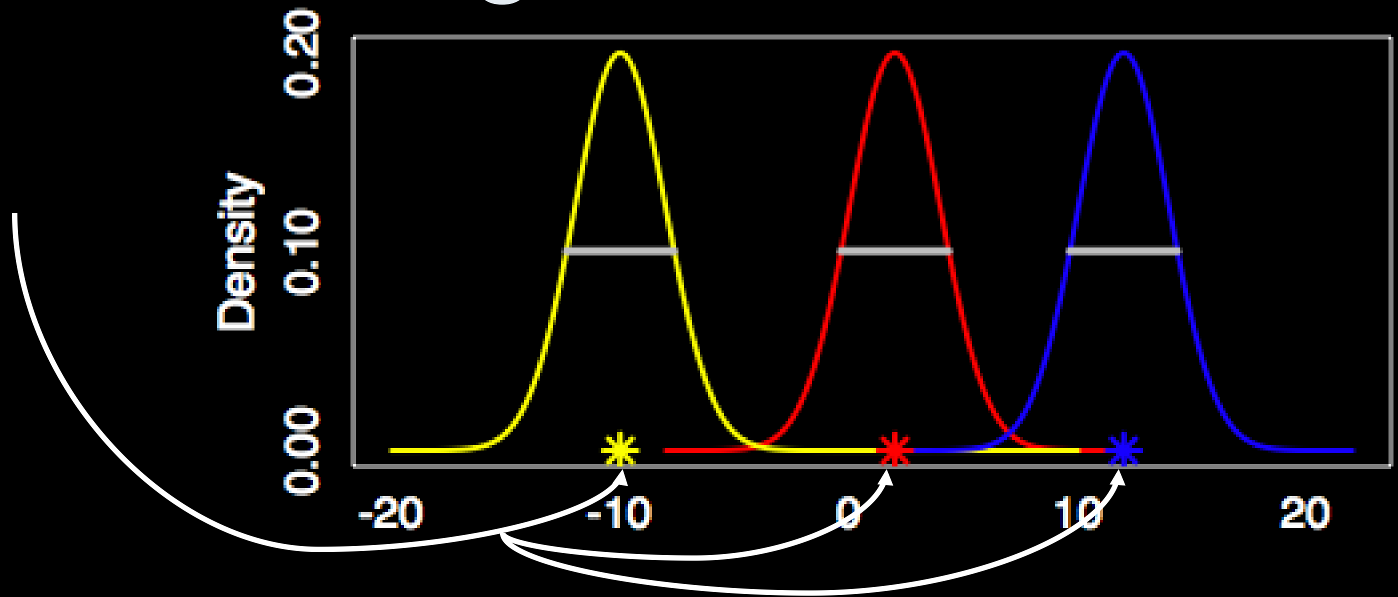
Stage 2

within between

$$\begin{pmatrix} \hat{\beta}_1 \\ \hat{\beta}_2 \\ \hat{\beta}_3 \end{pmatrix} = \begin{pmatrix} 1 \\ 1 \\ 1 \end{pmatrix} \beta_g + \begin{pmatrix} \eta^*_1 \\ \eta^*_2 \\ \eta^*_3 \end{pmatrix}, \quad \mathrm{Var}(\eta^*_i) = \frac{\sigma^2_{win}}{W} + \sigma^2_{btwn}$$

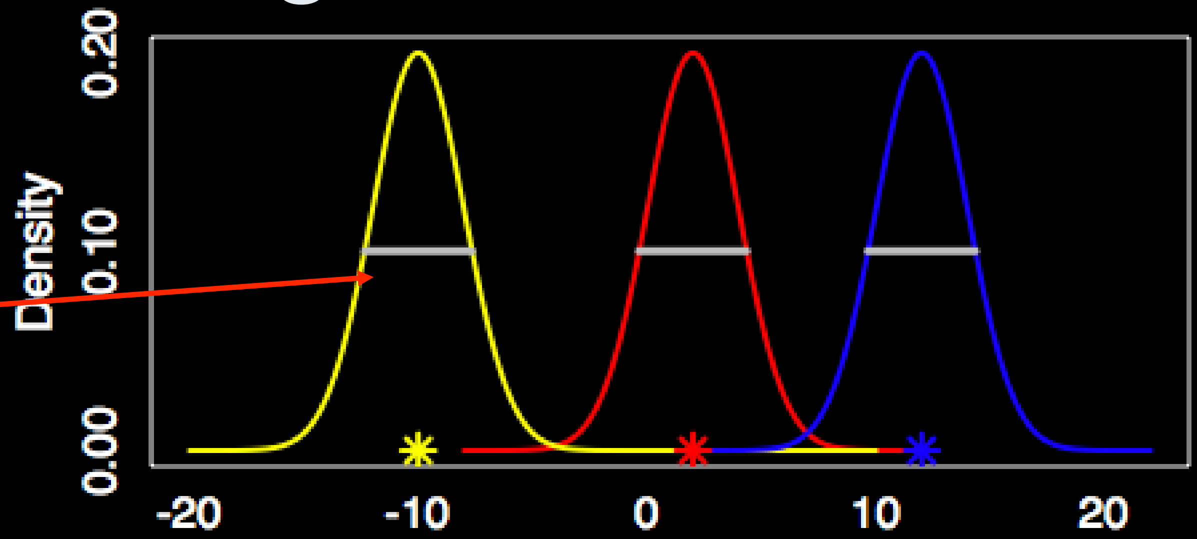
Two-Stage Model

- Stage 1
  - Means



Two-Stage Model

- Stage 1
  - Means
  - σ^2_{win} (same across subjects here)



Two-Stage Model

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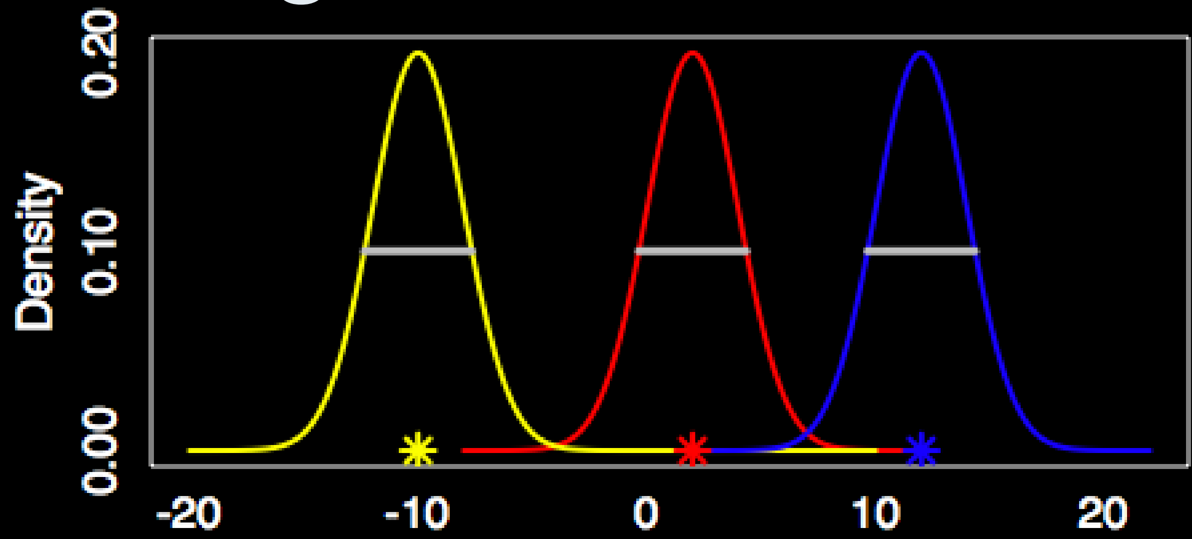
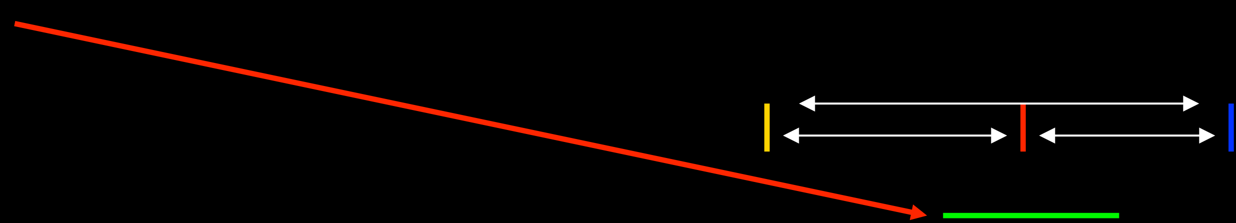
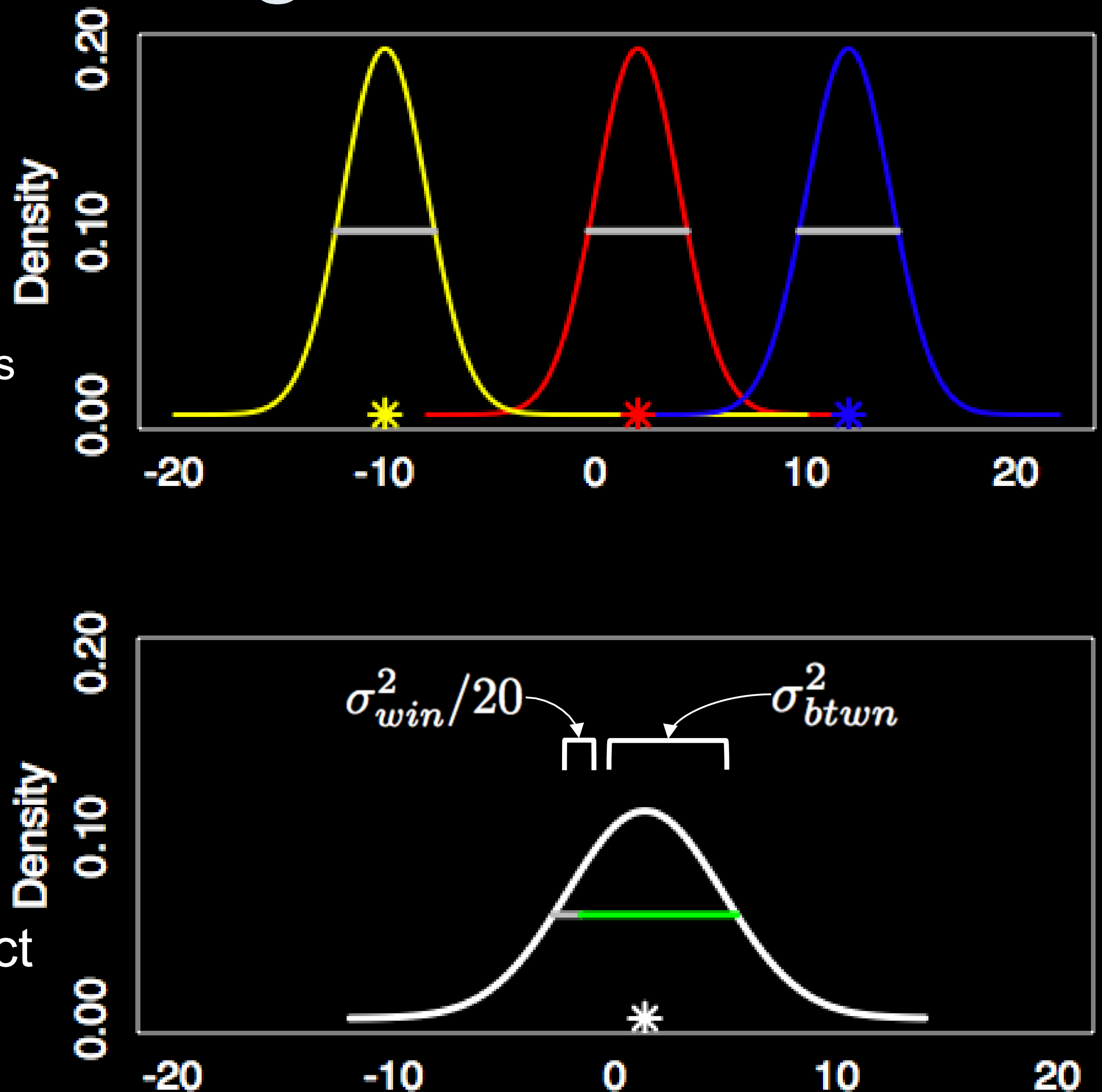


- Stage 2
  - σ^2_{btwn}



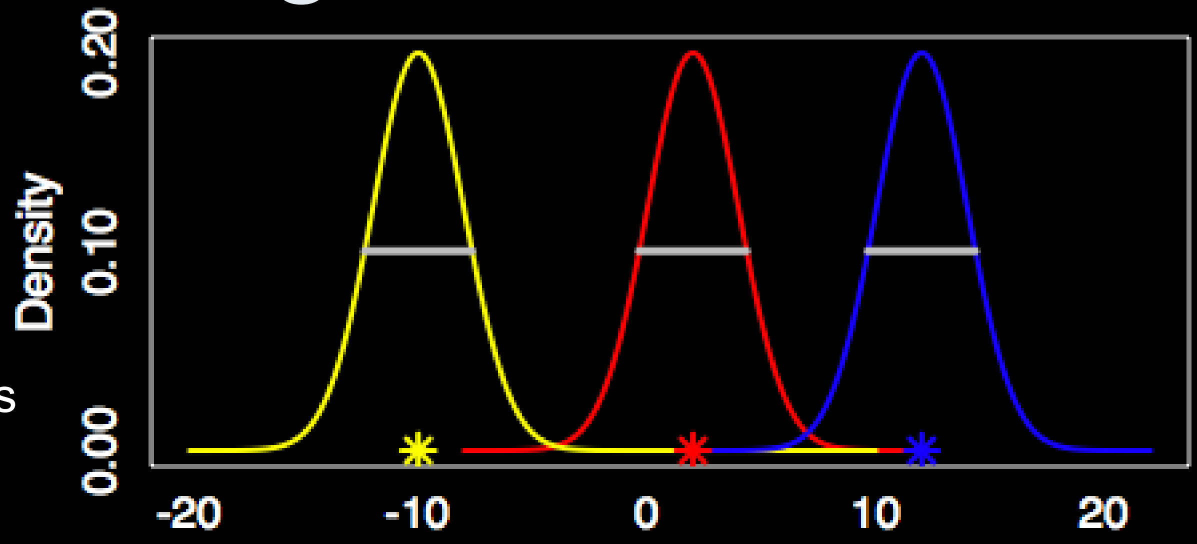
Two-Stage Model

- Stage 1
  - Means
  - σ^2_{win}
    (same across subjects here)
- Stage 2
  - σ^2_{btwn}
  - $\sigma^2_{mix} = \frac{\sigma^2_{win}}{20} + \sigma^2_{btwn}$
    - 20 hairs/subject

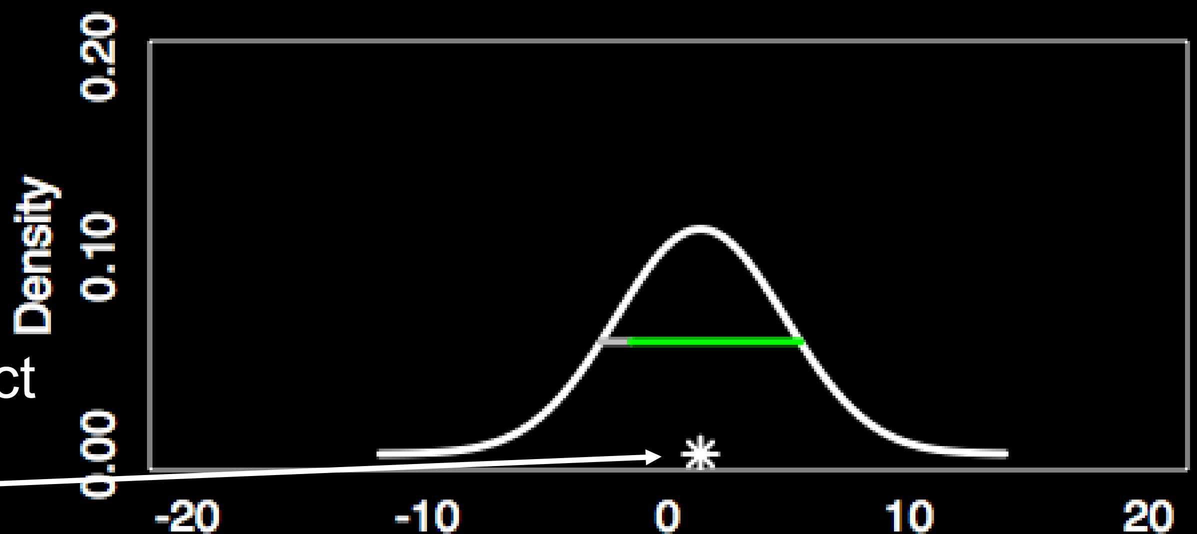


Two-Stage Model

- Stage 1
  - Means
  - σ^2_{win} (same across subjects here)



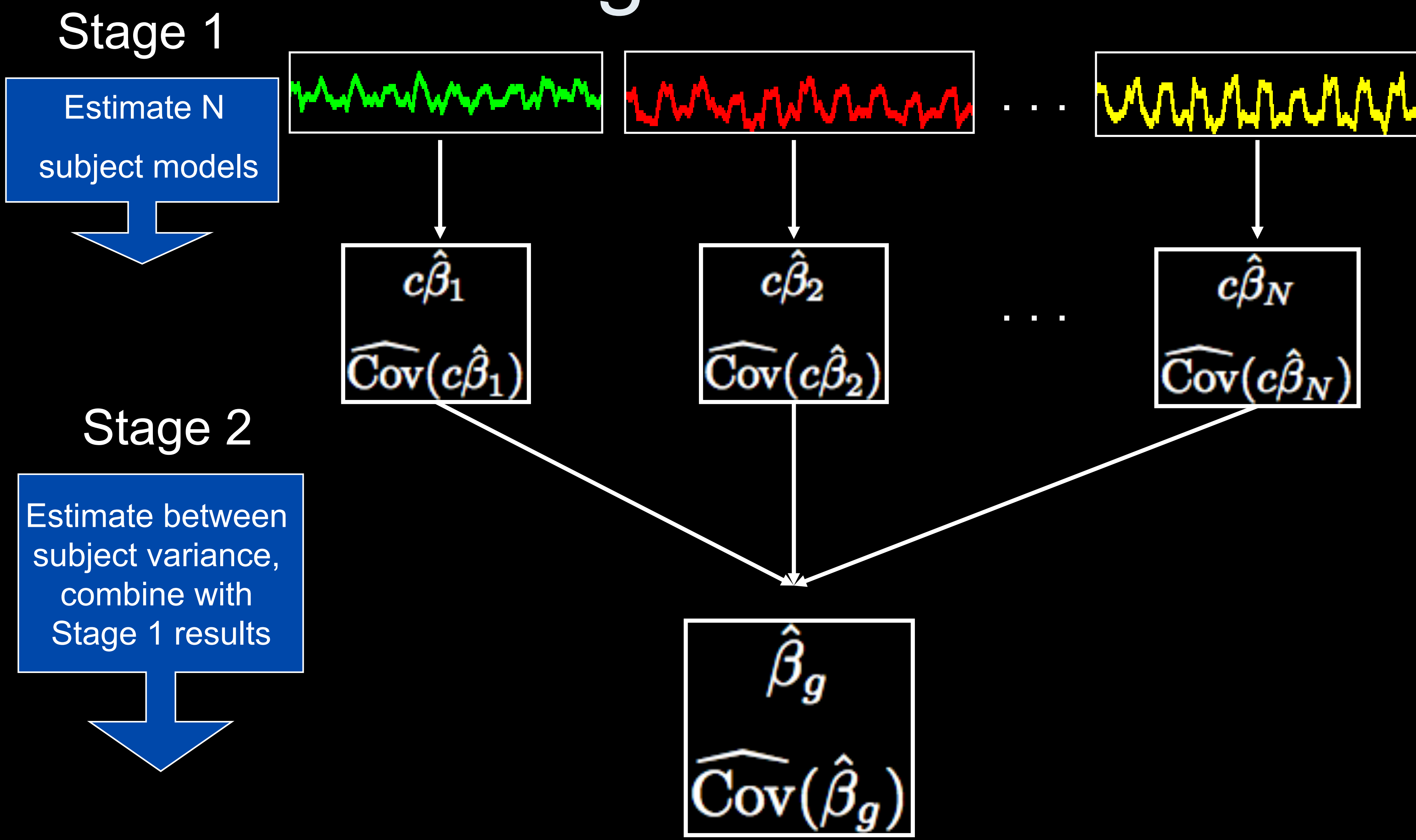
- Stage 2
  - σ^2_{btwn}
  - $\sigma^2_{mix} = \frac{\sigma^2_{win}}{20} + \sigma^2_{btwn}$
    - 20 hairs/subject
- Pop mean



Two-Stage Model

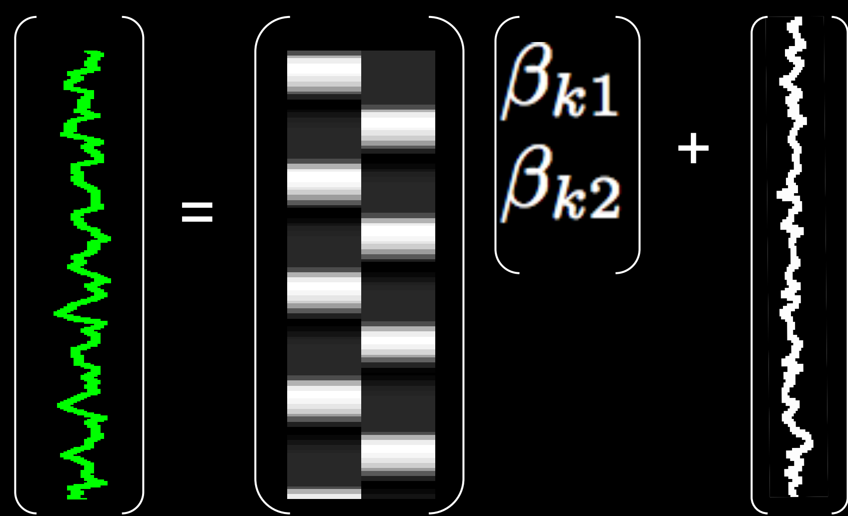
- $$T = \frac{\sqrt{N}\hat{\beta}}{\sqrt{\sigma^2_{win}/W + \sigma^2_{btwn}}}$$
  - N = # subjects
  - W = # measures within subject
- If new data are added, only run first stage for new data

Two Stage Model fMRI



Stage 1: Subject Model

$$Y_k = X_k \beta_k + \epsilon_k$$



$$\mathrm{Cov}(\epsilon_k) = \sigma_k^2 V_k$$

$$H_0 : \beta_{k1} - \beta_{k2} = 0$$

Stage 1: Estimation

- W_k such that $W_k V_k W_k' = I_T$

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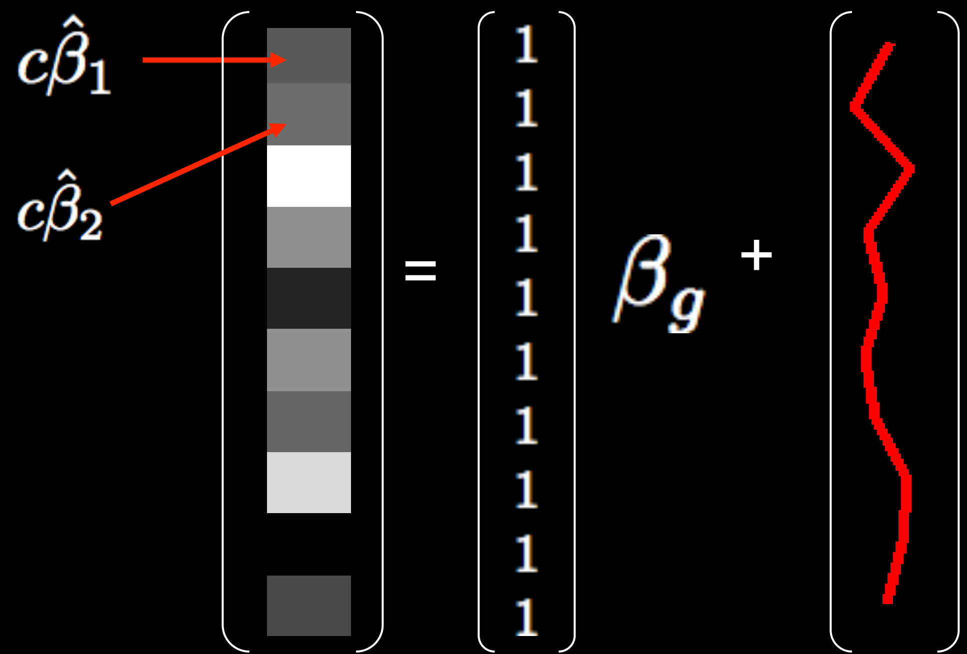
- W_k such that $W_k V_k W_k' = I_T$
- Whitened model
  - $W_k Y_k = W_k X_k \beta_k + W_k \epsilon_k$
  - $Y_k^* = X_k^* \beta_k + \epsilon_k^*$

Stage 1: Estimation

- W_k such that $W_k V_k W_k' = I_T$
- Whitened model
  - $W_k Y_k = W_k X_k \beta_k + W_k \epsilon_k$
  - $Y_k^* = X_k^* \beta_k + \epsilon_k^*$
- Use OLS on whitened model
  - $c\hat{\beta}_k = \left(X_k^{*\prime} X_k^* \right)^{-1} X_k^{*\prime} Y_k^*$
  - $\widehat{Cov}(c\hat{\beta}_k) = \hat{\sigma}_k^2 \left(X_k^{*\prime} X_k^* \right)^{-1}$

Stage 2: Group Model

$$\hat{\beta}_{cont} = X_g \beta_g + \epsilon_g$$



$$\text{Cov}(\epsilon_g) = V_g = \begin{pmatrix} \sigma_1^2 c(X_1^{*'} X_1^*)^{-1} c' & & \\ & \ddots & \\ & & \sigma_N^2 c(X_N^{*'} X_N^*)^{-1} c' \end{pmatrix} + \sigma_g^2 I_N$$

Stage 2: Estimation

- W_g such that $W_g V_g W_g' = I_N$

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- W_g such that $W_g V_g W_g' = I_N$
- $W_g \hat{\beta}_{cont} = W_g X_g \beta_g + W_g \epsilon_g$
  $\hat{\beta}^*_{cont} = X^*_g \beta_g + \epsilon^*_g$

Stage 2: Estimation

- W_g such that $W_g V_g W_g' = I_N$
- $W_g \hat{\beta}_{cont} = W_g X_g \beta_g + W_g \epsilon_g$
  $\hat{\beta}^*_{cont} = X^*_g \beta_g + \epsilon^*_g$

- $\hat{\beta}_g = \left(X^{*'}_g X^*_g \right)^{-1} X^{*'}_g \hat{\beta}^*_{cont}$
  $\widehat{Cov}(\hat{\beta}_g) = \left(X^{*'}_g X^*_g \right)^{-1}$

Stage 2: Estimation

- W_g such that $W_g V_g W_g' = I_N$
- $W_g \hat{\beta}_{cont} = W_g X_g \beta_g + W_g \epsilon_g$
  $\hat{\beta}^*_{cont} = X^*_g \beta_g + \epsilon^*_g$
- $\hat{\beta}_g = \left(X^{*\prime}_g X^*_g \right)^{-1} X^{*\prime}_g \hat{\beta}^*_{cont}$
  $\widehat{Cov}(\hat{\beta}_g) = \left(X^{*\prime}_g X^*_g \right)^{-1}$
- $T = \hat{\beta}_g / \sqrt{\widehat{\mathrm{Cov}}(\hat{\beta}_g)}$

Overview so far…

- The mixed model can be thought of as a two stage process
  - Stats software estimates both stages simultaneously
  - In fMRI it is computationally easier and more convenient to keep the stages separate
    - Easier to add new subjects
    - Not *exactly* the same as the stats software results

Overview so far

- When the model is estimated how is a subject with a high mfx variance treated differently than a subject with a low mfx variance?

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- When the model is estimated how is a subject with a high mfx variance treated differently than a subject with a low mfx variance?
  - Subjects with higher variability are down-weighted in the analysis!

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- What is a mixed effects model
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  - Random effects
- 2-stage summary statistics approach
- How do different software packages work?
- Overview FSL modeling options

How is the model estimated?

- Depends on software
  - SPM: Does not estimate σ_g^2
    - Due to a set of assumptions, estimation of is unnecessary
  - FSL: Bayesian approach to estimating σ_g^2

SPM2

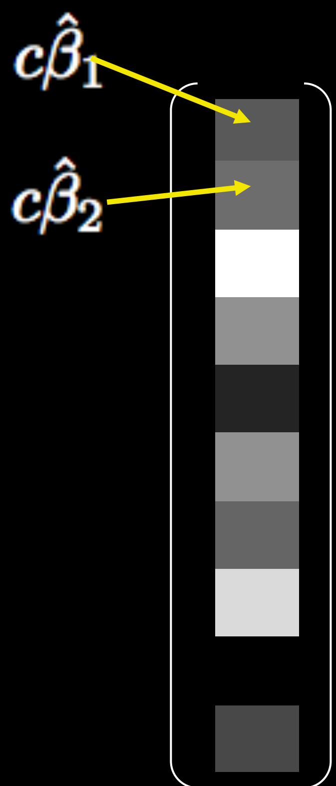
- Does not estimate σ_g^2
  - Assumes homogeneous variance across subjects
  - Assumes first level design is same across subjects

$$\hat{\sigma}^2_{win_{all}} = \hat{\sigma}^2_1 c \left(X_1^{*'} X_1^*\right)^{-1} c' = \ldots = \hat{\sigma}^2_N c \left(X_N^{*'} X_N^*\right)^{-1} c'$$

$$V_g = \sigma^2_{win_{all}} I_N + \sigma^2_g I_N = \sigma^2_{g*} I_N$$

OLS can be used

SPM2 : Single contrast per subject

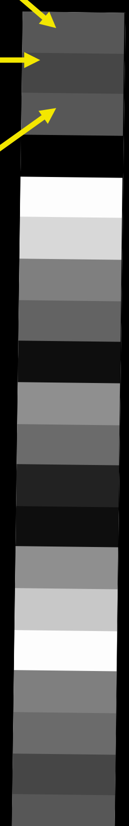


$$= \begin{pmatrix} 1 \\ 1 \\ 1 \\ 1 \\ 1 \\ 1 \\ 1 \\ 1 \\ 1 \\ 1 \end{pmatrix} \beta_g + \epsilon_g \qquad \epsilon_g \sim N(0, \sigma^2_{g*} I_N)$$

A one-sample T-test!

SPM2 : Multiple contrasts per subject

$\hat{\beta}_{1,1}$, $\hat{\beta}_{1,2}$, $\hat{\beta}_{2,1}$



$$
\left(\cdots\right) = \begin{pmatrix} 1 & 0 \\ 0 & 1 \\ 1 & 0 \\ 0 & 1 \\ 1 & 0 \\ 0 & 1 \\ 1 & 0 \\ 0 & 1 \\ 1 & 0 \\ 0 & 1 \\ 1 & 0 \\ 0 & 1 \\ 1 & 0 \\ 0 & 1 \\ 1 & 0 \\ 0 & 1 \\ 1 & 0 \\ 0 & 1 \\ 1 & 0 \\ 0 & 1 \end{pmatrix} \begin{pmatrix} \beta_{g_1} \\ \beta_{g_2} \end{pmatrix} + \epsilon_g
$$

$$
\epsilon_g \sim N(0, \sigma^2_{g*} \underline{\underline{V_{g^*}}})
$$

Global correlation estimate

SPM2 : Summary

- Multiple contrasts per subject can enter second level
  - Contrasts can be correlated
  - T and F-tests are possible
- Special case
  - One contrast per subject…Reduces to T-test!

SPM2

- Pros
  - Model is easy to estimate
  - Model is easy to understand
  - Multiple contrasts can enter the group model and are *not* considered independent
- Cons
  - Global covariance estimate (same across voxels)
  - Assumes variance is homogeneous across subjects

FSL: FMRIB Software Library

- Bayesian approach to estimating model
- Inference is based on *posterior* distribution of the data
  - $P(\beta_g, \sigma_g^2, \nu_g | Y)$
  - Parameters of interest are treated as random

FSL : Second Level Estimation

- Flame 1: Maximum a posteriori (MAP) estimate of σ_g^2 found iteratively
  - Assumes degrees of freedom, $\nu_g = N - p$
- Flame 2: Slower MCMC method of estimation
  - Applied to voxels close to threshold in step 1
  - Fine tunes estimates of $\beta_g, \sigma_g^2, \nu_g$
- Details
  - Woolrich et al. NI (2004) 1732-47

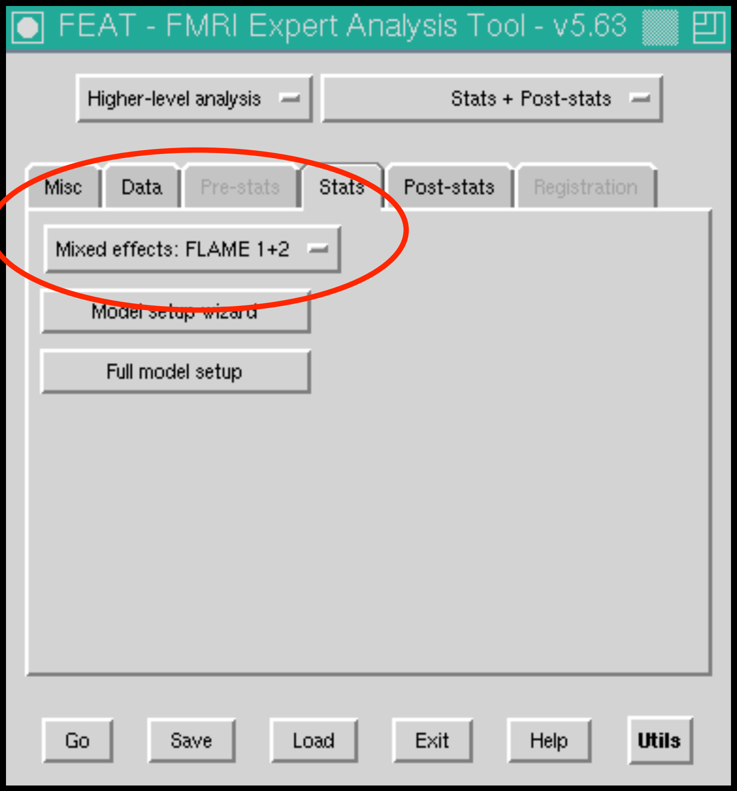
FSL

- Pros
  - When single contrast is taken to the second level, equivalent to all-in-one model
  - Within-subject variances are carried to the second level
    - Heterogeneity across subjects is modeled
- Cons
  - Multiple contrasts in the group model are assumed to be independent

Which software?

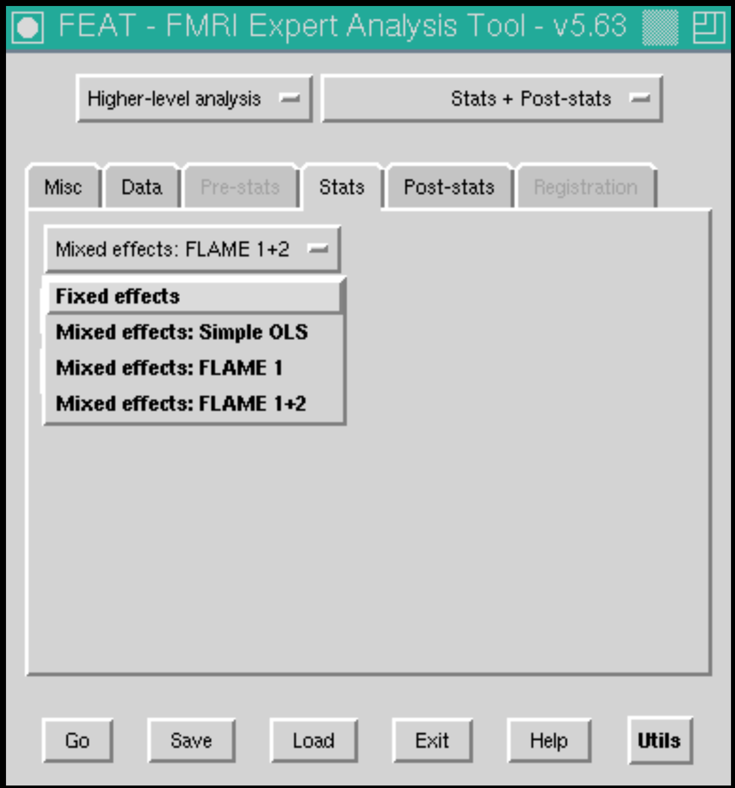
- FSL and FMRIstat best for heteroscedastic variances
  - Different number of trials per subject
- SPM best for multiple correlated contrasts at group level
- Other differences in first level modeling may sway users one way or another

FSL Group Model Options



FSL Group Model Options

- Fixed effects
  - Only uses w/in sub variance
- Simple OLS
  - Assumes w/in sub variances are equal
- Flame 1& 2
  - w/in sub var and btwn sub var



W_g Matrix

- Recall we pre-multiply by W_g so our errors are uncorrelated and constant variance

$$W_g \hat{\beta}_{cont} = W_g X_g \beta_g + W_g \epsilon_g$$

Act as weights

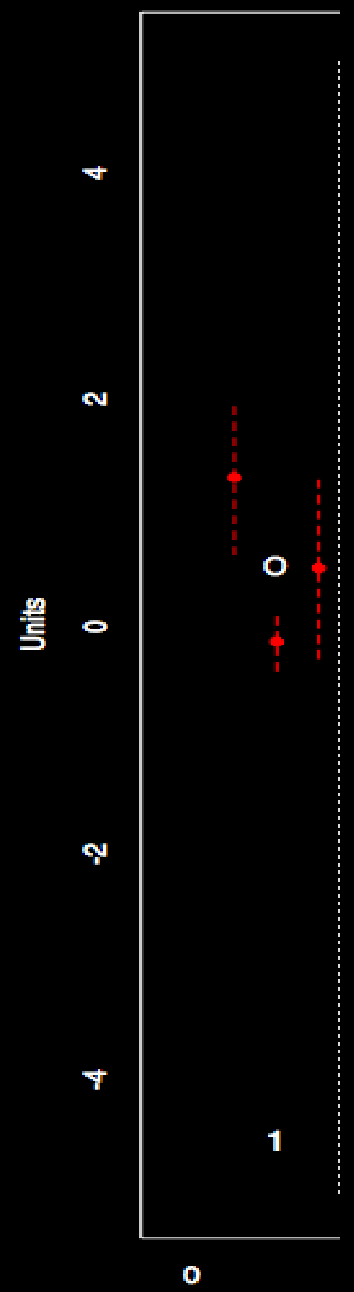
$$V_g = \begin{pmatrix} \sigma^2_{win_1} + \sigma^2_g & & 0 \\ & \ddots & \\ 0 & & \sigma^2_{win_N} + \sigma^2_g \end{pmatrix} \rightarrow W_g = \begin{pmatrix} \frac{1}{\sqrt{\sigma^2_{win_1} + \sigma^2_g}} & & 0 \\ & \ddots & \\ 0 & & \frac{1}{\sqrt{\sigma^2_{win_N} + \sigma^2_g}} \end{pmatrix}$$

V_g Matrix Assumptions

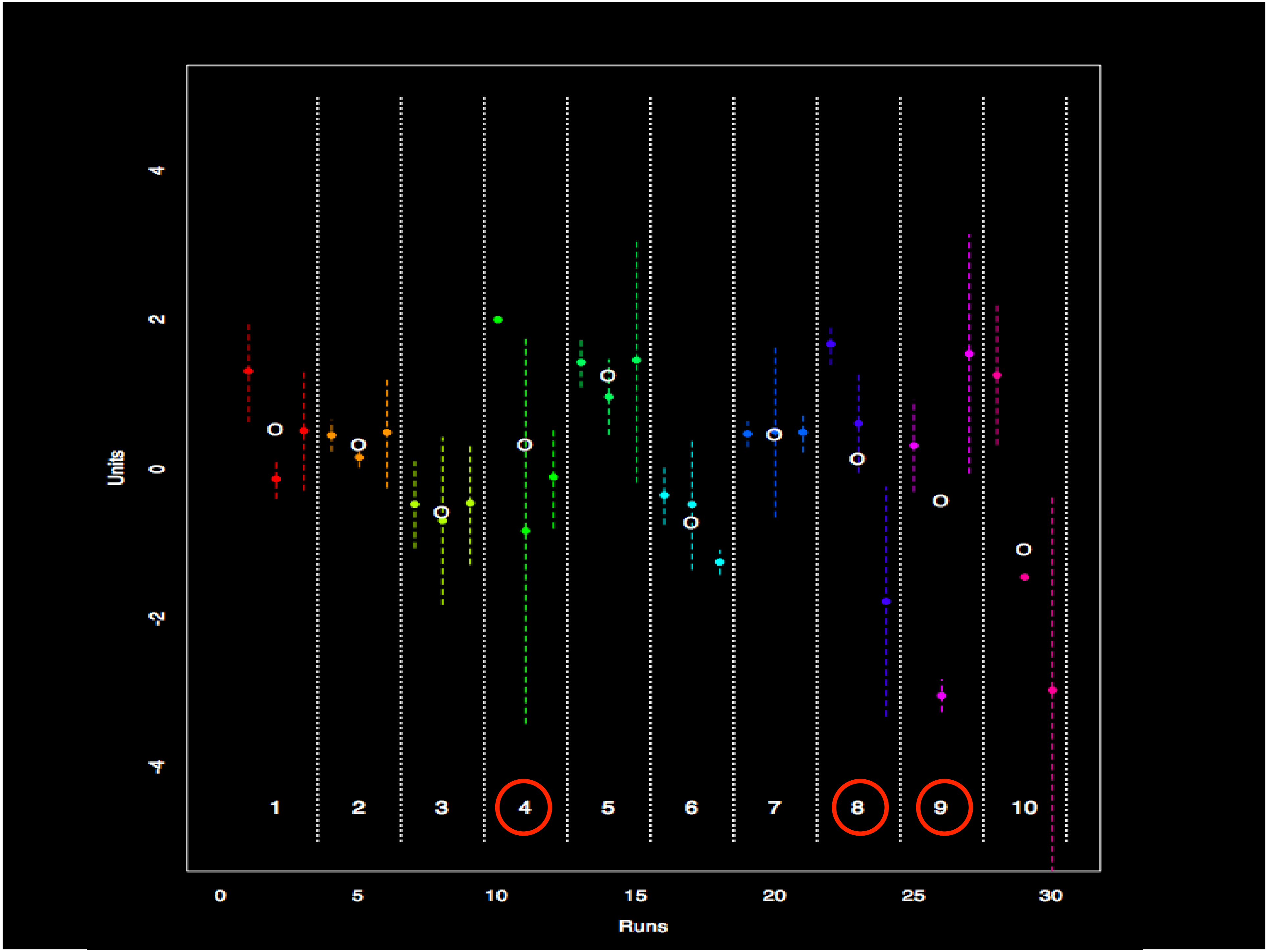
- Fixed effects analysis
  - Only appropriate for intermediate levels
  - Assumes the between-run variability=0

$$V_g = \begin{pmatrix} \sigma^2_{win_1} & & 0 \\ & \ddots & \\ 0 & & \sigma^2_{win_N} \end{pmatrix}$$

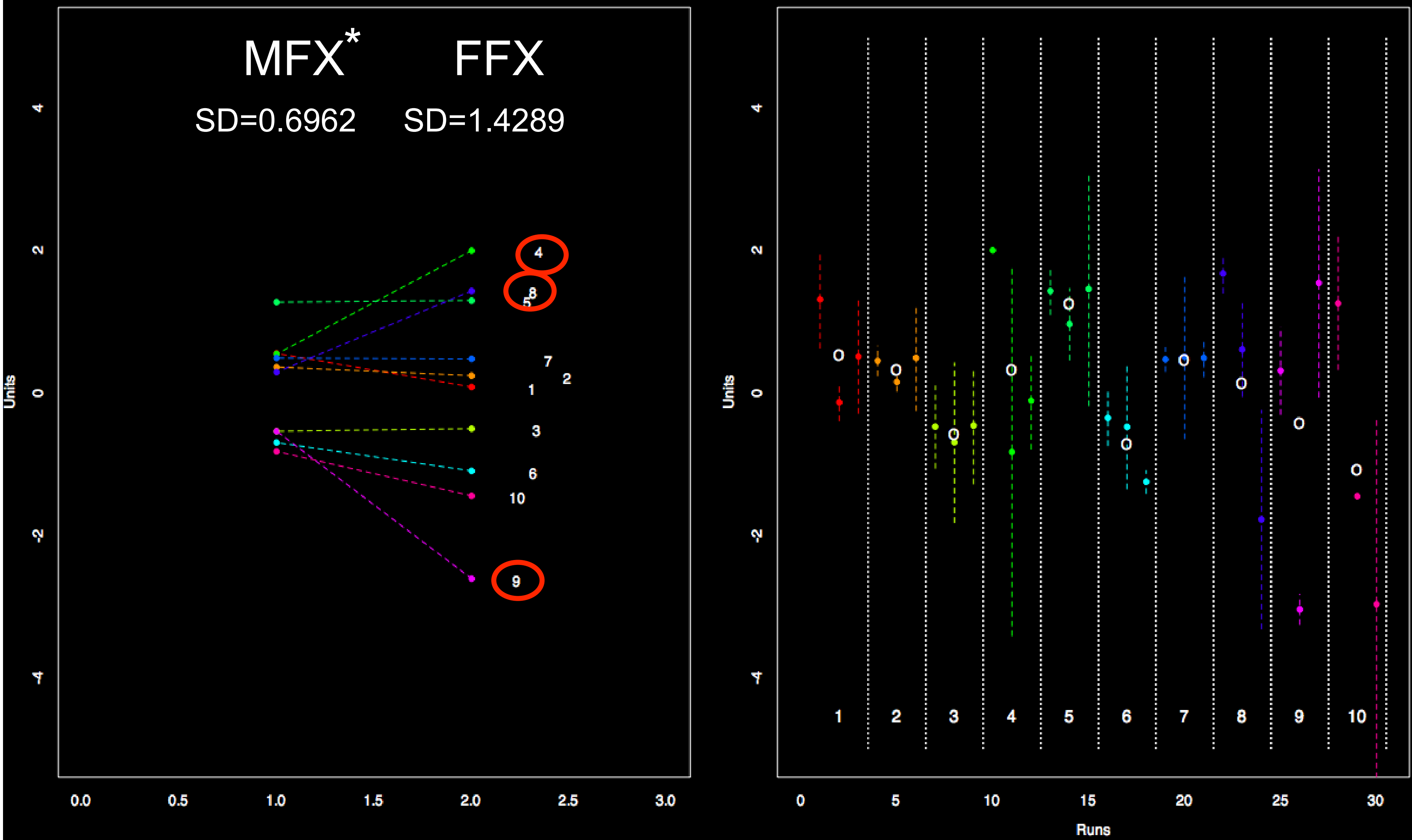
- Why would we do this?
  - What if df are low?
    - $\hat{\sigma}^2_g$ has high variance
    - If $\hat{\sigma}^2_g$ is too large, it will override differences in $\hat{\sigma}^2_{win_k}$



- 3 contrast estimates from 3 runs for 1 subject
- Dotted line indicates variance
- "o" marks unweighted mean



Units
4
2
0
-2
-4
1
2
3
4
5
6
7
8
9
10
0
5
10
15
20
25
30
Runs



*Assuming overestimate of $\hat{\sigma}_g^2$

Fixed Effects

- Use to improve your mean estimates
  - eg correct trials
- Since variance is underestimated, you *must* only run this at an intermediate level
  - Higher level analysis soaks up rest of variance

Third Level Analysis

- Typically Flame and OLS have similar results
  - Flame is probably the best choice, since it adjusts for heterogeneous variance
  - OLS runs faster
  - OLS stats can be larger or smaller than Flame stats
- FE at level 3 is bad
  - Variance is underestimated
  - High risk of false positives

Concluding Remarks

- Mixed models are appropriate for fMRI data
  - Include between-subject variance
  - Allows inference to be applied to entire population
- The two-stage summary statistics model
  - Computationally easier to estimate
  - Easier to add new subjects
- Software packages use the same basic model, but estimate σ_g^2 differently
- Use FE at intermediate levels and Flame at the top level in FSL

Shameless plug

Coming soon to an Amazon near you…